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# **Research Article**

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Eco-concern; Reliability; Scale; Sustainable nutrition; Validity

#### **Abbreviations:**

CFA, confirmatory factor analysis; EFA, exploratory factor analysis; EREC, Eating-Related Eco-Concern questionnaire; ICC, intraclass correlation coefficient

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# Turkish adaptation of the Eating-Related Eco-Concern questionnaire: validity and reliability study

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#### **Abstract**

The study aimed to translate the Eating-Related Eco-Concern (EREC) questionnaire into Turkish, adapt it cross-culturally, and evaluate its psychometric properties. EREC is a ten-item scale measuring how adults consider ecological impact in food choices due to climate change concerns. The study was conducted in Mersin between November 2023 and February 2024 with 442 adults (18-65 years) through face-to-face interviews. The Turkish version was adapted using the Translation-Back Translation method, and language validity was ensured. Face validity was evaluated through a pilot study with forty participants. Construct validity was initially assessed using exploratory factor analysis (EFA) with data from 200 participants, followed by confirmatory factor analysis (CFA) conducted on an independent sample of 242 participants to verify the factor structure. The reliability of the scale was assessed by test-retest analysis with 106 participants from the main sample (n 442), and consistency was measured by the intraclass correlation coefficient (ICC). The factor structure and model fit were evaluated using indices such as Comparative Fit Index (CFI), Goodness-of-Fit Index (GFI) and root mean square error of approximation (RMSEA). 68-3 % of the participants were female, whereas 31-7 % were male. The group's median BMI value was 24·3 (21·6–27·1), with 52·9 % classified as 'normal' according to the BMI classification. Factor 2 had a low score, but it was found to be adequate for other factors and the total scale score. The Turkish adaptation of the EREC questionnaire has been found to be a valid and reliable scale, as confirmed by comprehensive evaluations.

Climate change has become the most important threat to human beings in the twenty-first century<sup>(1)</sup>. A recent UN report on climate change reported that the level of global warming has been negatively affecting human beings<sup>(2)</sup>. Climate change, food systems and food security are interrelated terms. The food consumed and the production methods of food affect human health<sup>(3)</sup>. Reports on the causes and consequences of climate change have highlighted the importance of food systems and diets for the planet<sup>(4–7)</sup>. Ecological nutrition refers to the dietary pattern aimed at protecting human health and the planet's ecosystem. The transition to plant-based diets is thought to provide various environmental benefits, including reduced pressure on soil and water resources, less pollution and lower nitrogen and carbon-based greenhouse gas emissions. Ecological concerns arising from dwindling food resources have been reported to affect individuals' dietary preferences, leading to behavioural changes such as limiting meat consumption and avoiding food waste<sup>(8,9)</sup>. Additionally, many dietary changes required to reduce the global burden of disease can also contribute to a decrease in greenhouse gas emissions<sup>(10)</sup>. Therefore, due to environmental factors and climate change, individuals may wish to modify their eating habits<sup>(11–13)</sup>.

Climate change and ecological crises also affect the mental health of people by triggering their emotional distress<sup>(14)</sup>. As a result, the term eco-anxiety has emerged to describe the anxiety experienced in response to the ecological crisis<sup>(15)</sup>.

The Climate Change Worry Scale is another scale that is used to investigate the correlation between environmental concerns and food-related behaviours<sup>(16)</sup>. The Eating-Related Eco-Concern (EREC) questionnaire is a ten-item scale that assesses the extent to which individuals consider ecological impact when making food choices due to concerns about a changing climate. Its items are based on clinical observations and psychometric properties of the previously published Climate Change Worry Scale and a literature review on eco-friendly eating and sustainable nutrition. Items of EREC are rated on a 5-point Likert scale with 1 = 'Never', 2 = 'Rarely', 3 = 'Sometimes', 4 = 'Often' and 5 = 'Always'. The scale score is obtained by summing the scores given to the items and ranges from 10 to 50<sup>(17)</sup>. Higher scores signify that the level of eating-related eco-concern also elevates. The EREC questionnaire was initially developed in the USA, and then a validity and reliability study of the questionnaire was conducted in Lebanon. A study conducted on this questionnaire in Lebanon aimed to assess the psychometric properties of the scale for use in different cultural contexts<sup>(18)</sup>. Afterwards, a validity and reliability study was conducted in Italy, examining the cultural adaptability and



measurement sensitivity of the EREC<sup>(19)</sup>. In the literature, the relationship between ecological anxiety and eating disorders was evaluated using the EREC questionnaire in the study 'The Role of Eco-Anxiety in the Presentation of Bulimia Nervosa: A Case Report'<sup>(20)</sup>. Although there are some scales in Turkiye that have been validated and are reliable to assess the relationship between climate change, anxiety and eco-anxiety<sup>(21-23)</sup>, to the best of our knowledge, there is no valid and reliable scale to assess food-related eco-concern and nutritional relationships.

Age, BMI and gender are among the important variables affecting individuals' eco-concern levels. It is widely observed that younger individuals report higher levels of climate change and ecoconcern compared with older individuals (24-26). This may be explained by the fact that young people are more sensitive to the future impacts of the climate crisis and are more likely to directly experience its long-term consequences. On the other hand, BMI was found to be negatively correlated with eco-concern, which may be related to the greater adoption of restrictive dietary patterns by individuals with lower BMI(27,28). In terms of gender, it has been found that women are more concerned about climate change, have higher risk perceptions and are more likely to believe that climate change is happening compared with men<sup>(17,27,28)</sup>. However, a study on the EREC questionnaire has not shown a significant difference between genders, suggesting that food-related eco-concerns may have different dynamics than general climate concerns.

The aim of this study is to translate and cross-culturally adapt the widely used EREC questionnaire to screen the level of ecoconcern and changes in eating-related behaviours<sup>(17–19)</sup> into Turkish and to examine the psychometric properties and reliability and validity of the scale.

#### Method

## Research plan and sample selection

The study was conducted between November 2023 and February 2024 with the participants who were aged between 18–65 years, were living in Mersin and volunteered to participate in the study.

Data were collected through face-to-face interviews. In scale studies, it is recommended to include 5–10 participants for each scale item. Accordingly, a sample group of 200 participants was formed for exploratory factor analysis (EFA), which is twenty times the number of items in the ten-item EREC questionnaire. The number of participants used in confirmatory factor analysis (CFA) should not be lower than the number used in EFA. Then, CFA was conducted on a different sample group of 242 participants. Adequate sample size was achieved<sup>(29–31)</sup>.

Inclusion criteria were determined as follows: being aged between 18–65, having no communication problems and being voluntary to participate in the study. Exclusion criteria were determined as follows: having communication problems, being pregnant or breastfeeding, being illiterate and being foreign nationals. In this study, the EREC questionnaire was administered to adult individuals and individuals between the ages of 18–65 in accordance with the age classification of the WHO were included in the study<sup>(32)</sup>. In addition, pregnant and breastfeeding individuals were not included in the study because the physiological changes that occur during pregnancy have an impact on weight gain, eating habits, physical activity and psychological state. Changes in body image and eating behaviours during pregnancy may lead to increased vulnerability to eating disorders. Social factors such as media, family and friends have also been reported to be effective in

this process<sup>(33–36)</sup>. Before starting the study, permission was obtained from Baiyu Qi, one of the scale authors, via e-mail. The ethics committee approval was obtained from Nuh Naci Yazgan University Scientific Research and Publication Ethics Committee on 23/10/2023 with the decision numbered 2023/009-012 to conduct the study. The participants were informed about the study, and their written and verbal consent was obtained. A sociodemographic data form had sixteen questions about the participants' gender, age, marital status, educational level, family type, dietary habits, chronic diseases and medications. Anthropometric measurements (five questions) and body composition of the participants were taken before the study. BMI was calculated by dividing body weight in kilograms by the square of height in metres. The BMI was evaluated based on the classification provided by the WHO. According to this classification, individuals with a BMI of < 18.50 kg/m<sup>2</sup> are classified as underweight, those with a BMI between 18·50 and 24·99 kg/m<sup>2</sup> as normal, individuals with a BMI between 25.00 and 29.99 kg/m<sup>2</sup> as overweight and those with a BMI > 30.00 as obese. In the present study, the participants were divided into two groups: those with a BMI below 25.00 kg/m<sup>2</sup> (underweight and normal) and those with a BMI of 25.00 kg/m<sup>2</sup> or above (overweight and obese)(37). In the analyses, BMI was used as a categorical independent variable, not as a continuous covariate. The participants were grouped based on the BMI criteria of the WHO, and the comparisons were made between the two BMI categories:  $< 25.00 \text{ kg/m}^2 \text{ and } \ge 25.00 \text{ kg/m}^2.$ 

The Translation–Back Translation method was performed. After the necessary revisions were made, a pilot study was conducted, and the questionnaire was applied to the sample. The EREC questionnaire, consisting of ten questions, was applied by the face-to-face interview method.

# Turkish adaptation procedures

The validity of the scale was assessed by evaluating language, content, face and construct validities. On the other hand, the reliability of the scale was evaluated by evaluating the time invariance and internal consistency analyses. While conducting the study, steps on the scale adaptation guide of the WHO were taken into consideration<sup>(38)</sup>. The language validity of the scale was ensured by using the Translation–Back Translation method. The English version of the scale presented by the authors was translated into Turkish by two academicians (T.K., İ.K.) who have a good command of English. The scale form was back-translated from Turkish to English by two independent translators (N.Y., J.A.D.C.) who have a good command of Turkish and compared with its original English version. The translated scale was independently evaluated by four academicians (E.B., G.G.Ş., N.K., S.Ç.) specialised in nutrition and dietetics.

The Davis technique was used to question translation suitability and content validity; index values and content validity ratio were calculated<sup>(39)</sup>. Experts gave the scale items a score between 1 and 4: 1 point = not appropriate; 2 points = somewhat appropriate (major revision required); 3 points = fairly appropriate (minor revision required); and 4 points = very appropriate. The content validity ratio value for each item was calculated by dividing the number of experts who gave 3 or 4 points to the scale items by the total number of experts. The content validity index value was calculated by summing the content validity ratio values of all items and dividing by the total number of items. It was considered sufficient in terms of content validity that the scale's content validity ratio value was greater than 0.70 and the content validity

index value was greater than 0.80. Necessary corrections were made in line with the expert opinions. In addition, the items were evaluated in terms of conceptual equivalence and cultural adaptation, taking into account the Turkish cultural structure and target group, and necessary changes were made. After this step, the scale was made suitable for a pilot study<sup>(40-42)</sup>.

#### Pilot study

A pilot study was conducted to evaluate the face validity of the Turkish version of the EREC. Face validity is a process of obtaining expert opinions on whether the items in the scale are directed towards the related subject. Statistical evaluations are not made in this study<sup>(43)</sup>. It has been reported that the pilot study should be applied to at least 30–40 people representing the sample<sup>(38)</sup>. In the cognitive interview applied to these forty participants, the participants were asked to read the scale questions aloud and explain each item in their own words. By this method, the comprehensibility, language simplicity and conceptual clarity of the scale items were evaluated. The participants' opinions and suggestions were carefully recorded, and all feedback was analysed.

#### Test-retest

The test-retest method was used to assess time invariance. According to the COSMIN guidelines, a sample size of over 100 participants is considered sufficient for retest measurements<sup>(41)</sup>. Based on this recommendation, the retest was conducted with 106 participants.

The scale was re-administered to 106 participants 2 weeks after the first study, and the consistency between the first study and the test-retest was examined. Test-retest reliability coefficients were evaluated using the intraclass correlation coefficient (ICC) based on a one-way random effects model in which participants were randomised. The ICC value is a statistical measure used to assess reliability and consistency between measurements. It is often used to determine the reliability of scales, tests or inter-observer assessments. There is a generally accepted scale for interpreting the ICC value. Accordingly, ICC values less than 0.5 indicate poor reliability, values between 0.5 and 0.75 indicate moderate reliability, values between 0.75 and 0.9 indicate good reliability and ICC values greater than 0.9 indicate excellent reliability(44). This classification helps to understand how reliable the measurement or assessment is and plays a critical role in assessing the scientific validity of the results. In order to increase reliability, it is recommended to standardise the measurement conditions, ensure inter-observer consistency and use an adequate sample size. The reliability level of the scale was determined by calculating Cronbach's  $\alpha$  coefficient<sup>(44)</sup>.

# Statistical analysis

Qualitative data were summarised with numbers and percentages, while quantitative data were summarised with mean and standard deviation. If the variance assumptions were met, group comparisons were made with an independent two-sample t test for two groups and one-way ANOVA for more than two groups. If the variance assumptions were not met, the Welch's t test and Welch's ANOVA test were used. Additionally, the Bonferroni adjustment method was used to solve the multiplicity problem.

For the reliability and validity study, item analyses and EFA were conducted with a sample of 200 participants, followed by CFA with an independent sample of 242 participants to verify the factor structure identified through EFA. Moreover, Cronbach's  $\alpha$ 

coefficient was calculated to assess internal consistency, and ICC was calculated to assess consistency in repeated measures (test-retest). In item analyses, item-remainder correlations and Cronbach's  $\alpha$  coefficient statistics were examined when the item was deleted. EFA was based on the polychoric correlation coefficient, and Bartlett's test of sphericity was used to assess the factorability of the data, and the Kaiser-Meyer-Olkin statistic was used to assess sampling adequacy. While principal axis factoring was used as the factor extraction method, the Promax method was used as the rotation method. In order to determine the number of factors, the number of eigenvalues greater than 1 (Kaiser criterion) and the parallel analysis method were evaluated. The rate of variance explained was taken into consideration and reported. The CFA was performed with the unweighted least squares method. These fit indices are crucial in assessing the validity and reliability of the model. The ratio of  $\chi^2$  to df is commonly used to evaluate the results of a  $\chi^2$  test. This ratio helps to identify how well the test results fit or match the expected values (45). Comparative Fit Index, Goodness-of-Fit Index, non-normed fit index and Tucker-Lewis Index were used to evaluate the goodness of fit. The Bentler-Bonett's Normed Fit Index, root mean square error of approximation and standardised root mean square residual statistics were calculated. Results were reported as coefficient estimates, SE of the coefficient estimates, coefficient CI and p-values for the coefficients, as well as standardised coefficients and  $R^2$  statistics.

The ratio of  $\chi^2$  to df less than 2 indicates a good fit, while a value less than 3 is considered an acceptable fit (46,47). Goodness-of-Fit Index represents the overall fit of the model, values of > 0.95indicate a good fit and values of > 0.90 represent an acceptable fit (48). Similarly, the Comparative Fit Index assesses the model fit in comparison to an alternative model, with > 0.95 indicating a good fit and > 0.90 indicating an acceptable fit (48,49). Tucker-Lewis Index accounts for model complexity in its assessment. A value of > 0.95 indicates a good fit, while a value > 0.90 represents an acceptable fit(48,50). Normed Fit Index measures the model's improvement over the baseline model, where > 0.95 represents a good fit and > 0.90 represents an acceptable fit (51). Root mean square error of approximation indicates a good fit when values are low, with < 0.05 representing a good fit and < 0.08 representing an acceptable fit<sup>(52)</sup>. Standardised root mean square residual measures the difference between observed and predicted correlations, with values < 0.05 indicating a good fit and < 0.08 indicating an acceptable fit (48). The  $R^2$  represents the ratio of variance explained by the model, where a value > 0.10 is considered acceptable<sup>(53)</sup>. The Type I error level was set at 5 %. Analyses were performed using JASP software (Version 0.18.3, University of Amsterdam).

## Result

The majority of the participants were female (68·3 %) and were aged between 18 and 44 years (92·5 %). Most of the participants were single (50·4 %) and were living with their nuclear families (82·6 %). 75·8 % of the participants had a bachelor's degree or higher. 65·4 % were employed, and 47·5 % had an income equal to their expenses. According to the BMI classification, 52·9 % of the participants were normal weight, and 34·2 % were mildly overweight.

# Validity

To assess content validity, Content Validty Ratito (CVR) values were calculated as 1.00 for all items, resulting in a Content Validity Index (CVI) value of 1.00. Given that a CSR value above 0.70 and a

**Table 1.** Item analyses and results of exploratory factor analysis (n 200)

		Item analysis					
		SD	Item- reminder correlation	Cronbach's α when item deleted	кмо	Factor loads	
F1: Factors affecting food choices							
M9: I avoid foods that come with excess or non-recyclable packaging.	3.095	1.154	0.594	0-763	0.843	0.86	
M8: I try to only eat organic foods or food produced without pesticides.	2.935	1.061	0-657	0-744	0.835	0.819	
M10: I pay close attention to information on the impact that certain foods have on the environment (e.g., overfishing, greenhouse gasses, irrigation).	2.905	1.128	0-624	0-753	0.892	0.754	
M7: I avoid GM foods due to concerns about biodiversity loss.	2.920	1.183	0.601	0.761	0.801	0.564	
M1: I spend more time than other people searching for sustainable food.	2-610	0.986	0.463	0-800	0.927	0.406	
F1 Total	14-465	4.132		CA = 0.803	VE = 0·24		
F2: Food preferences due to climate change concern							
M3: I avoid eating any animal products due to my concerns about climate change.	1.540	0.838	0.661	0.580	0.702	0.961	
M2: I avoid eating meat due to concerns about climate change.	1.555	0.866	0.617	0-619	0·6 <sup>67</sup>	0.943	
M6: I try to eat less because of my concerns about climate change.	2.050	1.164	0-490	0-808	0.879	0.555	
F2 Total	5.145	2.360		CA = 0.744	VE = 0·210		
F3: Behavioural change due to climate change concern							
M4: I try not to waste food due to concerns about climate change.	3.565	1.298	0.510	_	0.794	0.844	
M5: I actively encourage others to change their behaviours to slow climate change.	2.765	1.228	0-510	-	0-844	0.73	
F3 Total	6-330	2.194		CA = 0.674	VE =	0.144	
Total scale	25-940	6.883		CA = 0.827	VE =	0.602	

 $\bar{x}$ : mean; CA, Cronbach's  $\alpha$ ; VE, variance explained; KMO, Keiser-Meyer-Olkin; EFA, exploratory factor analysis.

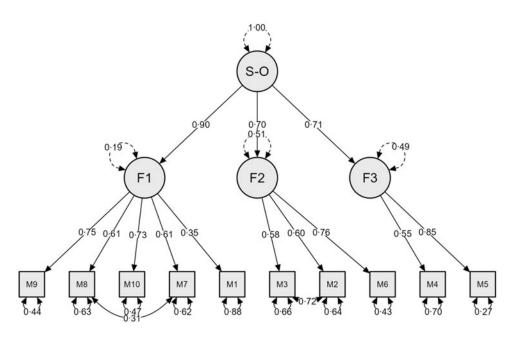
CGI value above 0.80 indicate sufficient content validity, the scale was deemed adequate in this regard. In the pilot study conducted with forty participants, it was observed that all scale items were clear and understandable, with no indications of ambiguity or confusion. Consequently, no revisions were necessitated, and the original version of the scale was retained. The data of the pilot study were not included in the main analysis. Construct validity was evaluated through a two-step statistical approach, beginning with EFA to identify the underlying factor structure, followed by CFA to validate and confirm the model fit. Item-rest correlation analysis showed that factor loadings ranged from 0.490 to 0.661, while the overall factor loadings ranged from 0.406 to 0.860. Cronbach's  $\alpha$  coefficients were calculated as 80·3 % for Factor 1 (M9, M8, M10, M7 and M1), 74·4 % for Factor 2 (M2, M3 and M6), 67.4 % for Factor 3 (M4 and M5) and 82.7 % for the overall scale. Following the evaluation of the modification indices, the suggested modifications between M2 and M3 (referring to avoiding meat and animal products due to concerns about climate change) and between M7 and M8 (referring to avoiding GM foods due to concerns about biodiversity loss and trying to eat only organic or pesticide-free foods) were implemented. These modifications resulted in a decrease of 16.49 and 14.17, respectively, in the total  $\chi^2$  value, indicating an improved model fit.

As a result of EFA, it was observed that the total explained variance ratio was 60.2 %. Although item 6 was removed from the

scale, Cronbach's  $\alpha$  value increased for the factor related to the removal of the item; it was not preferred to reduce the number of items to two, considering that Cronbach's  $\alpha$  statistic was sufficient. Since the lowest Kaiser–Meyer–Olkin statistic was 0·667 and Bartlett's test of sphericity was significant (P < 0.001), the factorability of the data and sampling adequacy were ensured (Table 1). In summary, there was a three-factor structure in which all statistics were at a sufficient level. This adequately fitting structure was then tested using CFA on an independent sample. After considering the modification suggestions, covariance was defined between M7 and M8 and between M2 and M3. Statistically significant correlations were found between these items (Fig. 1, standardised coefficients were calculated as 0·31 and 0·72, respectively).

All path coefficients obtained as a result of CFA were statistically significant (P < 0.001). When the statistics regarding the fit of the model were examined, it was found that the  $\chi^2/\text{SD}$  (1.353) statistic was less than 2; Goodness-of-Fit Index (0.995), Comparative Fit Index (0.995), Tucker–Lewis Index (0.993) and Normed Fit Index (0.983) statistics were greater than 95.0 %; and root mean square error of approximation (0.038) and standardised root mean square residual (0.044) statistics were less than 5.0 %. The smallest  $R^2$  statistic was calculated as 12.3 % (Table 2). According to all indices, it can be asserted that the model fit the data or the data validated the model.

<sup>\*</sup>Factor extraction method: principal axis factoring. Rotation method: Promax. Correlation matrix used: polychoric. Bartlett's test:  $\chi^2 = 926.57$ , df = 45, P < 0.001, root mean square error of approximation = 0.088, standardised root mean square residual = 0.024, Comparative Fit Index = 0.988, Tucker-Lewis Index = 0.92.



**Figure 1.** Model coefficients for second-level confirmatory factor analysis.

## Reliability

The test was administered again with 106 participants who were included in the CFA, and the test-retest consistency was examined with ICC statistics. Although Factor 2 showed a relatively low value (ICC = 0.648), the other factors and the total scale score were at a sufficient level (> 0.70) (Table 3). The model obtained through EFA was tested with CFA and found to be valid. When each item analysis, EFA, CFA and test-retest results were evaluated together, it was observed that the scale was valid and reliable, especially in terms of the use of the total scale score, and all goodness-of-fit statistics were adequate.

#### **Group comparisons**

When the gender groups were compared in terms of the scale scores, it was found that the mean score of female participants was statistically significantly higher than the mean of men in all dimensions except for the Factors Influencing Food Choice factor and in the total scale score (Table 4, P < 0.05). When BMI groups were compared in terms of scale scores, the mean of the  $< 25 \text{ kg/m}^2$ BMI group was found to be statistically significantly higher than the mean of the  $\geq 25 \text{ kg/m}^2 \text{ BMI group in the factors of Food}$ Choices Due to Climate Change Concern and Behavioural Changes Due to Climate Change Concern (Table 4, P < 0.05). The participants with BMI values in the normal and underweight ranges were significantly more likely to change their food preferences and behavioural changes due to climate change anxiety than slightly obese and obese individuals. When income groups were compared in terms of scale scores, the difference between mean scores of the groups was not statistically significant (finding not in the Table, P > 0.05).

## **Discussion**

In this study, the EREC<sup>(17)</sup>, a short and validated screening tool to assess eco-concern levels and changes in eating-related behaviours, was translated into Turkish, and its validity and reliability were evaluated. While scales assessing anxiety on climate change are

available in Turkey<sup>(21,22)</sup>, there is no validated tool that comprehensively assesses food-related eco-concern<sup>(23)</sup>. The Turkish version of the EREC questionnaire is the third language adaptation of this scale. The original version of the scale was developed in the USA and adapted into Arabic and Italian before the present study<sup>(18,19)</sup>.

To examine the construct validity, the results of CFA showed that certain errors in covariates added to the one-factor model of the original scale improved the fit of the model to the data. Specifically, error terms were found to share common variance between the items that refer to avoiding meat and animal products due to concerns about climate change (M2 and M3) and the items that include avoiding GM foods due to concerns about biodiversity loss (M7) and trying to only eat organic or pesticide-free foods (M8). These covariations are theoretically and logically consistent although they are not reported in previous EREC validation studies. In particular, items 2 and 3 reflect dietary changes shaped by concern about the effects of animal food consumption on climate change. In the CFA, significant covariates between items 2 and 3 and items 7 and 8 indicated that climate change concern was associated with both avoidance of meat/animal product consumption and tendencies to consume organic and non-GM foods, which is compatible with previous studies (9,54-56). As a result, modelling error covariations between scale items strengthened the construct validity of the EREC and better explained the overlaps between dietary preferences based on environmental concerns. These covariations were also observed in the Italian adaptation of the original scale, supporting the consistency and validity of findings of the present study<sup>(19)</sup>.

In the original EREC questionnaire, factor analysis results (including parallel analysis, Kaiser's rule and scree plot) showed that the EREC questionnaire consisted of a single dimension<sup>(17)</sup>. In the Arabic adaptation study, the results of the parallel analysis, which accounted for 69·05 % of the variance, showed that the scale had a unidimensional structure as in the original scale<sup>(18)</sup>. In its Italian adaptation study, the results of CFA conducted to assess the internal construct validity of the scale showed that the model fit the data well with the addition of error covariates between specific items<sup>(19)</sup>. In the present study, the results of EFA indicated that there was a three-factor structure in which all statistics were at a

**Table 2.** Results of confirmatory factor analysis (n 242)

					95 CI%		
	Coefficient	SE	Z	Р	Lower, upper	Std. coefficient	R <sup>2</sup>
F1: Factors influencing food choices							
M9: I avoid foods that come with excess or non-recyclable packaging.	0-420	0.115	3-634	< 0.001	0.193, 0.646	0.750	0.56
M8: I try to only eat organic foods or food produced without pesticides.	0.315	0.088	3.588	< 0.001	0.143, 0.486	0-607	0.36
M10: I pay close attention to information on the impact that certain foods have on the environment (e.g. overfishing, greenhouse gasses, irrigation).	0.407	0-112	3.635	< 0.001	0-188, 0-627	0-726	0.52
M7: I avoid GM foods due to concerns about biodiversity loss.	0-357	0.099	3.601	< 0.001	0.162, 0.551	0.614	0.37
M1: I spend more time than other people searching for sustainable food.	0-173	0.049	3.541	< 0.001	0.077, 0.268	0-351	0.12
F2: Food choices due to climate change concern							
M3: I avoid eating any animal products due to my concerns about climate change.	0-348	0.054	6-444	< 0.001	0-242, 0-454	0-580	0.33
M2: I avoid eating meat due to concerns about climate change.	0.339	0.053	6-389	< 0.001	0.235, 0.444	0-603	0.36
M6: I try to eat less because of my concerns about climate change.	0.540	0.079	6-859	< 0.001	0.385, 0.694	0-757	0.57
F3: Behavioural changes due to climate change concern							
M4: I try not to waste food due to concerns about climate change.	0.560	0.058	9-601	< 0.001	0.446, 0.674	0.545	0.29
M5: I actively encourage others to change their behaviours to slow climate change.	0.734	0.083	8-889	< 0.001	0.572, 0.896	0-854	0.73
S-O: Second level							
FIFC	2.070	0.683	3.028	0-002	0.730, 3.409	0-900	0.81
FCDCCC	0-977	0.166	5.882	< 0.001	0.651, 1.303	0-699	0.48
BCDCCC	1.016	0.149	6-821	< 0.001	0.724, 1.308	0.713	0.50

FIFC, Factors Influencing Food Choice; FCDCCC, Food Choices Due to Climate Change Concern; BCDCCC, Behavioural Changes Due to Climate Change Concern. Std. coefficient. standardised coefficient.

 $Compliance \ statistics: \ \chi^2 = 40\cdot59, \ df = 30, \ \chi^2/df = 1\cdot353, \ Goodness-of-Fit\ Index = 0\cdot995, \ Comparative\ Fit\ Index = 0\cdot995, \ Tucker-Lewis\ Index = 0\cdot993, \ Normed\ Fit\ Index = 0\cdot983, \ root\ mean\ square\ error\ of\ approximation = 0\cdot038, \ standardised\ root\ mean\ square\ residual = 0\cdot044.$ 

Table 3. Consistency for test-retest (n 106)

	ICC
Factor 1	0.744
Factor 2	0.648
Factor 3	0.712
Total Scale	0.765

ICC, intraclass correlation coefficient.

sufficient level: F1: Factors Influencing Food Choices  $\rightarrow$  (M9, M8, M10, M7, M1); F2: Food Choices Due to Climate Change Concern  $\rightarrow$  (M3, M2, M6); and F3: Behavioural Changes Due to Climate Change Concern  $\rightarrow$  (M4, M5). This structure was further tested with another sample (participant group) by CFA. According to all analyses, it can be asserted that the model fit the data or the data confirmed the model. Unlike the original EREC questionnaire, rather than fully retaining the structure, the main themes of the scale remained similar, but the factor organisation changed. This difference may be due to the culture-specific interpretation of the scale, the different ways in which environmental concerns are perceived or differences in sample size.

In the EREC questionnaire, the Cronbach  $\alpha$  coefficient was found to be 0.87 in the subset<sup>(17)</sup>, 0.91 in the Arabic adaptation<sup>(18)</sup> and 0.90 in the Italian adaptation<sup>(19)</sup>. In the Turkish adaptation, this value is 0.82. This value was found to be sufficient (> 0.70) for the overall scale<sup>(44)</sup>.

Environmental anxiety is believed to drive individuals towards more sustainable eating habits and reduce the consumption of animal-based foods. In this context, it is suggested that environmental anxiety and concerns about climate change can shape food preferences and lead individuals to adopt more sustainable eating behaviours<sup>(55)</sup>. However, the literature shows that women experience higher levels of environmental anxiety than men, and this anxiety may be more effective in guiding women towards more sustainable eating habits. Women are found to carry higher levels of anxiety and a greater sense of responsibility regarding climate change<sup>(57)</sup> and are more willing to engage in environmentally friendly behaviours<sup>(58)</sup>. This has been confirmed in studies showing that women are more likely to adopt sustainable eating behaviours, such as increasing organic food consumption and reducing meat intake.

In the literature, some studies have reported no significant difference between the eco-anxiety levels of women and men<sup>(59,60)</sup>. However, other studies have indicated that male university students have lower environmental behaviour scores compared with their

Table 4. Comparison of scale scores according to gender and BMI groups

	Female ( <i>n</i> 302)								< 25 BMI (kg/m²) (n 248)		≥ 25 BMI (kg/m²) (n 194)				
	Mean	SD	Mean	SD	t	P*	Cohen d	Mean	SD	Mean	SD	t	P*	Cohen d	
FIFC	14-21	4.23	13.56	4.51	1.458	0.146	0.149	14	4.31	14.01	4.36	-0.003	0.998	0.000	
FCDCCC	5.17	2.42	4-22	1.82	4.564	$< 0.001^{\dagger}$	0.422	5.13	2.47	4.54	1.98	2.754	0·006 <sup>†</sup>	0-257	
BCDCCC	6-26	2.15	5.59	2.45	2.775	0·006 <sup>†</sup>	0.297	6.3	2.18	5.73	2.34	2.666	0.008	0-256	
Total scale	25-64	7.02	23-38	6-96	3.157	0.002	0.323	25-43	7.2	24-27	6-87	1.712	0.088	0.164	

FIFC, Factors Influencing Food Choice; FCDCCC, Food Choices Due to Climate Change Concern; BCDCCC, Behavioural Changes Due to Climate Change Concern. \*Independent two-sample t test. ¹Welch t test.

female counterparts(61). The literature indicates that women are more likely to believe that climate change is happening<sup>(62)</sup>, perceive more climate change risk<sup>(63)</sup> and are more concerned about it<sup>(64)</sup>. In a study examining the relationships between mental health, lifestyle factors and climate change concerns among Norwegian adolescents, it was found that female students were more likely to worry about climate change. However, no gender difference was observed in eco-anxiety related to eating habits (65). Likewise, another study assessing the eco-concern levels of university students and their association with eating habits revealed that female students had higher scale scores compared with male students. This finding indicated that female students had higher levels of eco-anxiety than male students. Additionally, when the relationship between the ecoconcern scale score and questions related to climate change was evaluated, it was reported that students with higher eco-concern levels believed that they had more individual influence on climate change<sup>(66)</sup>. Women were more concerned about environmental health and felt a stronger sense of personal responsibility. This may explain why environmentally sustainable dietary behaviours, such as increasing consumption of organic food<sup>(67)</sup> and reducing meat consumption<sup>(68)</sup>, are more common among women<sup>(69)</sup>.

In the original study of the EREC questionnaire, it was found that there was a significant difference between male and female participants in terms of climate change anxiety (t(221) = -2.90, P = 0.0042) and female participants showed higher levels of climate change anxiety compared with male participants (mean = 30.53, SD = 8.06; mean = 26.57, SD = 9.06 for men). However, this anxiety did not have any effect on eating behaviours<sup>(23)</sup>. In the Arabic adaptation of the scale, no significant difference was found between men and women in terms of scale scores (mean = 16.39, SD =  $5.93 \, v$ . mean = 15.98, SD = 5.98, t(761) = 0.905, P = 0.366)<sup>(18)</sup>. In the Italian adaptation, the sample was highly gender imbalanced, and 85% of participants were identified as female. As a result, no gender comparison was made in the Italian adaptation<sup>(19)</sup>.

The analyses showed that configurational, metric and scalar invariance was maintained between genders. This suggests that the scale measures similarly for both genders, and these findings are consistent with the initial EREC study. Although previous studies indicated that women had higher levels of general environmental concerns (e.g. about climate change), there were no gender differences in environmental concerns related to eating habits. This finding is compatible with the results of the Arabic adaptation of the scale. In this study, the majority of the participants were female (76·0 % in EFA and 62·0 % in CFA). When the gender groups were compared in terms of scale scores, it was found that the mean scores of female participants were statistically significantly higher than the mean scores of their male counterparts in all

factors except for the Factors Influencing Food Choice and in the total scale score (Table 4, P < 0.05). These findings suggested that women had higher levels of environmental concerns than men and that this also affected their food-related eco-concerns. There are also studies in the literature reporting that there is no significant difference in the level of eco-anxiety in women compared with men<sup>(59,60)</sup>. There is a need for further studies evaluating the relationships between eco-concern, gender and dietary habits in a more comprehensive manner.

There are studies in the literature suggesting that climate anxiety and eco-anxiety are more pronounced among younger individuals<sup>(24,25)</sup>. The mean age of the participants in the original EREC questionnaire and its Italian adaptation was 37 years (16,19), while in the Arabic adaptation, the mean age was 28 years (18). In both the original EREC questionnaire study and its Arabic adaptation study, the correlation between age and EREC scores has not been discussed. In its Italian adaptation study, no relationship was found between EREC scores and younger age<sup>(19)</sup>. This suggests that climate-related concerns are not exclusive to younger generations in the Italian context. One possible explanation for this difference could be the increasing frequency and severity of climate change events, such as extreme heat waves<sup>(70)</sup>. These events particularly lead to more pronounced health issues in older adults, suggesting that eco-anxiety affects not only young adults but also older individuals who are directly affected by such events<sup>(19)</sup>. In the present study, however, no clear relationship was established between age distribution and food-related eco-anxiety levels.

In recent years, the relationship between BMI, eco-concerns and dietary habits has attracted increasing attention. The related studies have shown that individuals' levels of eco-anxiety can affect their food choices and behavioural changes related to sustainability. For example, individuals with higher levels of eco-anxiety tend to adopt more environmentally friendly eating behaviours, such as reducing meat consumption and increasing the intake of plant-based and organic foods<sup>(9,54)</sup>. Furthermore, a relationship has been found between BMI and eco-concern levels, with some studies suggesting that normal-weight or underweight individuals are more likely to engage in pro-environmental dietary behaviours compared with those with higher BMI values<sup>(56)</sup>.

A study conducted to evaluate the eco-anxiety levels of university students and to examine the relationship between this anxiety and eating habits reported that there was no significant difference between BMI groups in terms of eco-anxiety scale scores  $(P > 0.05)^{(64)}$ . In the original scale study and its Italian adaptation study, it was found that the scale scores did not show a significant correlation with BMI and EREC<sup>(17,19)</sup>. There is no discussion about BMI in its Arabic adaptation study. However, when BMI groups

were compared in terms of scale scores in the present study, it was observed that the  $<25\cdot00~kg/m^2~BMI$  group scored statistically significantly higher on the Food Choices Due to Climate Change Concern and Behavioural Changes Due to Climate Change Concern factors compared with the  $\geq25\cdot00~kg/m^2~BMI$  group (Table 4,  $P<0\cdot05$ ). Normal-weight and underweight participants showed a significantly higher tendency to change their food preferences and alter their behaviours due to eco-concern compared with slightly overweight or obese individuals (Table 4). This may be attributed to the fact that individuals with low BMI tend to pay more attention to their health and environmental concerns lead them to change their behaviours more. On the other hand, this effect was less significant in obese or slightly overweight individuals, suggesting that the effect of environmental concerns on eating behaviours may vary depending on physical factors.

In conclusion, the EREC questionnaire, a ten-item scale that assesses the extent to which individuals consider the ecological impact when making food choices due to concerns about eco-anxiety, is a valid and reliable scale adapted to Turkish. This scale is used to evaluate the environmental concerns that shape individuals' dietary habits. The diversity of experiences and cultural contexts is of great importance in the application of assessment tools like the EREC questionnaire in different linguistic and cultural settings. This may pave the way for future studies and enable to conduct studies on eco-concern and dietary habits across various cultural contexts. It is recommended that this scale be used to evaluate environmental concerns and eating habits in more detail.

## Strengths and limitations

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The limitations of the study are that the study had a limited sample size; data were collected through self-reports, which may be affected by participant bias; and causal relationships were not examined due to the cross-sectional design. Additionally, while the validity and reliability of the Turkish version of the EREC questionnaire were tested, validation studies in different cultural contexts are lacking. The fact that the majority of the participants were aged between 18 and 44 also limits the evaluation of the effect of age. The age distribution reflects the group that voluntarily agreed to participate in the study, independent of the researchers. Thus, results cannot necessarily be generalised. In future studies using this scale, we recommend that the results be evaluated in terms of young adults by considering the limitations regarding the distribution of age groups. These limitations highlight the need for future research with larger and more diverse samples. In addition, a strength of the present study is that an acceptable sample size was used to examine the consistency and reliability of the EREC among Turkish adults. The EREC questionnaire is a viable and effective tool for assessing the level of eco-concern and changes in eatingrelated behaviours. The use of various factor analysis techniques, such as EFA and CFA, provides a robust methodological approach to verify the construct validity of the scale.

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## References

- Ripple W, Wolf JC, Newsome TM, et al. (2020) World scientists' warning of a climate emergency. Biosci 70, 8–12.
- Intergovernmental Panel on Climate Change (IPCC) (2023) Climate Change 2023: Synthesis Report. Contribution of Working Groups I, II and III to the Sixth Assessment Report of the Intergovernmental Panel on Climate Change. Geneva: IPCC. p. 184. –https://doi.org/10.59327/IPCC/ AR6-9789291691647
- 3. Schiavo M, Le Mouël C, Poux X, et al. (2021) An agroecological Europe by 2050: What impact on land use, trade and global food security? (Study No. 08/21). IDDRI. Retrieved August 2025, from https://www.iddri.org/sites/de fault/files/PDF/Publications/Catalogue%20Iddri/Etude/202107-ST0821\_TYFA%20World\_1.pdf
- Willett W, Rockström J, Loken B, et al. (2019) Food in the Anthropocene: the EAT-Lancet Commission on healthy diets from sustainable food systems. Lancet 393, 447–492.
- Searchinger T, Waite R, Hanson C, et al. (2019) Creating a Sustainable Food Future: A Menu of Solutions to Feed Nearly 10 Billion People by 2050. Washington, DC: World Resources Institute.
- Springmann M, Clark M, Mason D'Croz D, et al. (2018) Options for keeping the food system within environmental limits. Nature 562, 519–525.
- Watts N, Amann M, Arnell N, et al. (2019) The 2019 report of The Lancet Countdown on health and climate change: ensuring that the health of a child born today is not defined by a changing climate. Lancet 394, 1836–1878.
- Garza R (2023) Examining ecological harshness cues and anticipated food scarcity on food preferences. Food Qual Prefer 104, 104736.
- 9. Sanchez-Sabate R & Sabate J (2019) Consumer attitudes towards environmental concerns of meat consumption: a systematic review. *Int J Environ Res Public Health* **16**, 1220.
- Wang DD, Li Y, Afshin A, et al. (2019) Global improvement in dietary quality could lead to substantial reduction in premature death. J Nutr 149, 1065–1074.
- Kim BF, Santo RE, Scatterday AP, et al. (2020) Country-specific dietary shifts to mitigate climate and water crises. Glob Environ Change 62, 101926.
- Rose D, Heller MC, Willits-Smith AM, et al. (2019) Carbon footprint of self-selected US diets: nutritional, demographic, and behavioral correlates. Am J Clin Nutr 109, 526–534.
- Green RF, Joy EJM, Harris F, et al. (2018) Greenhouse gas emissions and water footprints of typical dietary patterns in India. Sci Total Environ 643, 1411–1418.
- Ojala M, Cunsolo A, Ogunbode CA, et al. (2021) Anxiety, worry, and grief in a time of environmental and climate crisis: a narrative review. Annu Rev Environ Resour 46. 35–58.
- Pihkala P (2020) Anxiety and the ecological crisis: an analysis of ecoanxiety and climate anxiety. Sustainability 12, 7836.
- Stewart AE (2013) Psychometric properties of the Climate Change Worry Scale. Int J Environ Res Public Health 18, 494.
- 17. Qi B, Presseller EK, Cooper GE, et al. (2022) Development and validation of an Eating-Related Eco-Concern Questionnaire. *Nutrients* 14, 4517.
- El Zouki CJ, Obeid S, Malaeb D, et al. (2024) Validation of the Arabiclanguage version of the eating-related eco concern (EREC) scale. PREPRINT (Version 1) available at Research Square. https://doi.org/10. 21203/rs.3.rs-3987459/v1
- 19. Tecuta L, Casu G & Tomba E (2024) Validation of the Italian version of the Eating-Related Eco-Concern Questionnaire: insights into its relationship with orthorexia nervosa. *Front Psychol* **15**, 1441561.
- Runfola CD & Safer DL (2024) The role of eco-anxiety in the presentation of bulimia nervosa: a case report. J Eat Disord 12, 164.

- Özbay S & Alcı B (2021) İklim Değişikliği Kaygı Ölçeği: Türkçeye Uyarlama, Geçerlik ve Güvenirlik Çalışması [Climate Change Anxiety Scale: Turkish Adaptation, Validity, and Reliability Study]. R&S Res Stud Anatolia J 4, 183–193.
- 22. Cebeci F, Karaman M, Öztürk AF, et al. (2022) İklim Değişikliği Anksiyetesi Ölçeği'nin Türkçe Uyarlaması: geçerlik ve Güvenirlik Çalışması [ The Turkish Adaptation of the Climate Change Anxiety Scale: Validity and Reliability Study]. Ufkun Ötesi Bilim Dergisi 22, 20–42.
- Uzun K, Öztürk AF, Karaman M, et al. (2022) Adaptation of the ecoanxiety scale to Turkish: a validity and reliability study. Arch Health Sci Res 9, 110–115.
- Clayton S & Karazsia BT (2020) Development and validation of a measure of climate change anxiety. J Environ Psychol 69, 101434.
- Rocchi G, Pileri J, Luciani F, et al. (2023) Insights into eco-anxiety in Italy: preliminary psychometric properties of the Italian version of the Hogg ecoanxiety scale, age and gender distribution. J Environ Psychol 92, 102180.
- Chung SJ, Jang SJ & Lee H (2023) Eco-anxiety and environmental sustainability interest: a secondary data analysis. *Int J Ment Health Nurs* 32, 1462–1472.
- Heeren A, Mouguiama-Daouda C & McNally RJ (2023) A network approach to climate change anxiety and its key related features. *J Anxiety Disord* 93, 102625.
- Reser JP & Bradley GL (2020) The nature, significance, and influence of perceived personal experience of climate change. WIREs Clim Change 11, e668.
- 29. Gorsuch RL (1974) Factor Analysis. Philadelphia: Saunders.
- Boateng GO, Neilands TB, Frongillo EA, et al. (2018) Best practices for developing and validating scales for health, social, and behavioral research: a primer. Front Public Health 6, 149.
- Floyd FJ & Widaman KF (1995) Factor analysis in the development and refinement of clinical assessment instruments. Psychol Assess 7, 286–299.
- World Health Organization (2001) Age Standardization Rates. GPE
  Discussion Paper Series 31 (Internet). Geneva: WHO. https://cdn.who.int/
  media/docs/default-source/gho-documents/global-health-estimates/gpe\_
  discussion\_paper\_series\_paper31\_2001\_age\_standardization\_rates.pdf
  (accessed 21 March 2025).
- 33. Fiat F, Merghes PE, Scurtu AD, et al. (2022) The main changes in pregnancy—therapeutic approach to musculoskeletal pain. *Medicina* **58**, 1115.
- Samra A & Dryer R (2024) Problematic social media use and psychological distress in pregnancy: the mediating role of social comparisons and body dissatisfaction. *Journal of Affective Disorders* 361, 702–711.
- Chauhan A & Potdar J (2022) Maternal mental health during pregnancy: a critical review. Cureus 14, e30656.
- 36. Saleem T, Saleem S, Shoib S, *et al.* (2022) A rare phenomenon of pregorexia in Pakistani women: need to understand the related behaviors. *J Eat Disord* **10**, 1–11.
- 37. Weir CB & Jan A (2024) BMI Classification Percentile and Cut off Points. Treasure Island, FL: StatPearls.
- Çapık C, Gözüm S & Aksayan S (2018) Kültürlerarası ölçek uyarlama aşamaları, dil ve kültür uygulaması: Güncellenmiş rehber [Cross-cultural scaling adaptation stages, language and culture application: Updated guide]. Florence Nightingale J Nurs 26, 199–210.
- Davis LL (1992) Instrument review: getting the most from a panel of experts. Appl Nurs Res 5, 194–197.
- Ayre C & Scally AJ (2014) Critical values for Lawshe's content validity ratio: revisiting the original methods of calculation. Meas Eval Couns Dev 47, 79–86.
- Lawshe CH (1975) A quantitative approach to content validity. Pers Psychol 28, 563–575.
- Wilson FR, Pan W & Schumsky DA (2012) Recalculation of the critical values for Lawshe's content validity ratio. *Meas Eval Couns Dev* 45, 197–210.
- Tavşancıl E (2002) Tutumların Ölçülmesi ve SPSS ile Veri Analizi [Measuring Attitudes and Data Analysis with SPSS], 1st ed. Ankara: Nobel Yayınevi.

 Koo TK & Li MY (2016) A guideline of selecting and reporting intraclass correlation coefficients for reliability research. J Chiropr Med 15, 155–163.

- 45. Field A (2013) Discovering Statistics using IBM SPSS Statistics, 4th ed. London: SAGE Publications.
- Kline RB (2016) Principles and Practice of Structural Equation Modeling, 4th ed. New York: Guilford Press.
- Schermelleh-Engel K, Moosbrugger H & Müller H (2003) Evaluating the fit of structural equation models: tests of significance and descriptive goodness-of-fit measures. Methods Psychol Res Online 8, 23–74.
- Hu LT & Bentler PM (1999) Cutoff criteria for fit indexes in covariance structure analysis: conventional criteria versus new alternatives. Struct Equ Model 6, 1–55.
- Bentler PM (1990) Comparative fit indexes in structural models. Psychol Bull 107, 238–246.
- Tucker LR & Lewis C (1973) A reliability coefficient for maximum likelihood factor analysis. *Psychometrika* 38, 1–10.
- Bentler PM & Bonett DG (1980) Significance tests and goodness of fit in the analysis of covariance structures. Psychol Bull 88, 588–606.
- Browne MW & Cudeck R (1993) Alternative ways of assessing model fit. In Testing Structural Equation Models, pp. 136–162 [KA Bollen, JS Long, editors]. Newbury Park, CA: Sage.
- 53. Hair JF, Black WC, Babin BJ, et al. (2010) Multivariate Data Analysis: A Global Perspective, 7th ed. Upper Saddle River, NJ: Pearson Education.
- Cheah I, Shimul AS, Liang J, et al. (2020) Drivers and barriers toward reducing meat consumption. Appetite 149, 104636.
- 55. Lai AE, Tirotto FA, Pagliaro S, *et al.* (2020) Two sides of the same coin: environmental and health concern pathways toward meat consumption. *Front Psychol* 11, 578582.
- Turnes A, Pereira P, Cid H, et al. (2023) Meat consumption and availability for its reduction by health and environmental concerns: a pilot study. Nutrients 15, 3080.
- López L, Martínez E & Fernández A (2020) Gender differences in environmental attitudes and sustainable consumption behavior. *Environ* Psychol 38, 100–113.
- White MA, Habib R & Jafry T (2021) Gendered responses to environmental change: a study of eco-anxiety in young adults. J Environ Behav 56, 205–223.
- Patrick R, Snell T, Gunasiri H, et al. (2023) Prevalence and determinants of mental health related to climate change in Australia. Aust N Z J Psychiatry 57, 710–724.
- Whitmarsh L, Player L, Jiongco A, et al. (2022) Climate anxiety: what predicts it and how is it related to climate action? J Environ Psychol 83, 101866.
- Weerasekara PC, Withanachchi CR, Ginigaddara G, et al. (2020) Food and nutrition-related knowledge, attitudes, and practices among reproductiveage women in marginalized areas in Sri Lanka. Int J Environ Res Public Health 17, 3985.
- Clayton SD, Pihkala P, Wray B, et al. (2023) Psychological and emotional responses to climate change among young people worldwide: differences associated with gender, age, and country. Sustainability 15, 3540.
- Ergun SJ, Karadeniz ZD & Rivas MF (2024) Climate change risk perception in Europe: country-level factors and gender differences. *Humanit Soc Sci* Commun 11, 1573
- Bush SS & Clayton A (2023) Facing change: gender and climate change attitudes worldwide. Am Polit Sci Rev 117, 591–608.
- Leonhardt M, Granrud M, Bonsaksen T, et al. (2022) Associations between mental health, lifestyle factors and worries about climate change in Norwegian adolescents. Int J Environ Res Public Health 19, 12826.
- 66. Oral FN & Durmuş H (2023) Üniversite öğrencilerinde eko-anksiyete düzeyi ve beslenme alışkanlıklarıyla ilişkisi: Erciyes Üniversitesi örneği [The level of eco-anxiety among university students and its relationship with eating habits: The case of Erciyes University]. Sosyal Sağlık Derg 3, 15–25.
- Gundala RR, Nawaz N, Boobalan K, et al. (2022) Does gender moderate the purchase intention of organic foods? Theory of reasoned action. Heliyon 8, e10456.

68. Ripabelli G, Mastronardi L, Tamburro M, et al. (2017) Food consumption and eating habits: a segmentation of university students from Central-South Italy. New Medit 16, 56.

- 69. Wollmar M, Post A & Sjöberg A (2024) Food choice, activity level, and carbon footprint: exploring potential for sustainable food consumption practices in young adults. *Front Nutr* 11, 1449054.
- 70. World Meteorological Organization & Copernicus Climate Change Service (2025) European State of the Climate: Extreme Events in Warmest Year on Record (Internet). Geneva & Bonn: WMO & C3S; 2025 Apr 15. https://wmo.int/news/media-centre/european-state-of-climate-extreme-events-warmest-year-record (accessed 19 June 2025).